



**EVALUATING THE GENERALIZED BUCHSTAB FUNCTION AND
REVISITING THE VARIANCE OF THE DISTRIBUTION OF THE
SMALLEST COMPONENTS OF COMBINATORIAL OBJECTS**

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Abstract

Let $n \geq 1$ and let X_n be the random variable representing the size of the smallest component of a random combinatorial object made of n elements. Combinatorial objects belong to parametric classes. This article focuses on the exp-log class with parameter $K = 1$ (permutations, derangements, polynomials over a finite field, etc.) and $K = 1/2$ (surjective maps, 2-regular graphs, etc.). The generalized Buchstab function Ω_K is essential in evaluating probabilistic and statistical quantities. For $K = 1$, it is known that $\text{Var}(X_n) = C(n + O(n^{-\epsilon}))$ for some $\epsilon > 0$ and sufficiently large n . We revisit the evaluation of $C = 1.3070\dots$ using different methods: analytic estimation using tools from complex analysis, numerical integration using Taylor expansions, and computation of the exact distributions for $n \leq 4000$ using the recursive nature of the counting problem. In general, for any K , the quantity $1/\Omega_K(x)$ for $x \geq 1$ is related to the asymptotic proportion of n -objects with large smallest components. We show how the coefficients of the Taylor expansion of $\Omega_K(x)$ for $[x] \leq x < [x] + 1$ depend on those for $[x] - 1 \leq x - 1 < [x]$. We use this family of coefficients to evaluate $\Omega_K(x)$.

1. Introduction

Let $n \geq 1$ and let X_n be the random variable representing the size of the smallest component of a random combinatorial object made of n elements. By a random combinatorial object, we mean a combinatorial object chosen uniformly at random among all possible combinatorial objects of size n . The cardinality of the support of X_n is, in principle, $n + 1$. Since the length of the smallest component obviously

cannot be between $\lfloor n/2 \rfloor + 1$ and $n - 1$ inclusively, the range of X_n is $1, 2, \dots, \lfloor n/2 \rfloor$ together with n . For some reasons that will become clear hereafter, we add zero probabilities to extend the range of X_n over all integers between 1 and n inclusively.

The theory about combinatorial objects and analytical methods required to understand many of the references in this paper is in [6]. Our results in Section 2 are valid for the class that contains permutations, derangements, and monic polynomials over finite fields, to name a few. Our result in Section 3 applies to all combinatorial objects in the exp-log class. We let readers consult [6] for the proper definitions of the exp-log class of combinatorial objects.

We can take the typical permutations or monic polynomials over finite fields. The latter receives special treatment in [9]. References [12] and [13] give local results about the probability distribution of X_n and asymptotic results about the k -th moment of X_n . One of our goals in this paper is to revisit some results concerning the second moment to compute the variance of X_n , denoted by $\text{Var}(X_n)$. We recall that, by definition,

$$\text{Var}(X_n) = \sum_{k=1}^n (k - \mathbf{E}(X_n))^2 \mathbf{P}\{X_n = k\} = \mathbf{E}(X_n^2) - (\mathbf{E}(X_n))^2, \tag{1}$$

where $\mathbf{P}\{X_n = k\}$ is the probability that X_n equals k , and $\mathbf{E}(X_n)$ is the expectation of X_n .

The k -th moment of X_n , $\mathbf{E}(X_n^k)$, is expressed as an integral involving the ordinary Buchstab function ω , which is defined over the real interval $[1, \infty)$ by

$$\omega(x) = \frac{1}{x} \quad \text{for } 1 \leq x \leq 2, \quad \text{and} \quad \frac{d(x\omega(x))}{dx} = \omega(x - 1) \quad \text{for } x \geq 2. \tag{2}$$

As mentioned in [12], the k -th moment of X_n involves the quantity $\int_1^\infty t^{-k}\omega(t)dt$. Besides the original paper by Buchstab [3] in which the function is defined and analyzed, numerous other papers discuss its various properties and applications, such as [2]. The book [15] contains many proofs of the properties of the Buchstab function.

Theorem 5 from [13] stipulates that

$$\text{Var}(X_n) = C(n + O(n^{-\epsilon})) \quad \text{for some } \epsilon > 0. \tag{3}$$

The constant C from (3) is given by

$$C = 2 \int_1^\infty \frac{\omega(t)}{t^2} dt. \tag{4}$$

Remark 1. In [1], [11], [12], and [13], the integration interval for (4) starts at 2. When computing the variance, the authors inadvertently forgot to add $3/4$, resulting from the integration over the interval $[1, 2)$. This mistake leads to confusion for some researchers; see [5].

Let S_n be the set of permutations on n elements, and let $S_{k,n} \subsetneq S_n$ be those permutations with smallest cycles of length k for $1 \leq k \leq n$. Denote the cardinality of $S_{k,n}$ by $s_{k,n}$. Let $c_k = (k - 1)!$ for $k \geq 1$, and let $[n/k] = 1$ if and only if $k|n$; otherwise $[n/k] = 0$. Then, in [12] it is proved that

$$s_{k,n} = \sum_{i=1}^{\lfloor n/k \rfloor} \frac{c_k^i}{i!} \frac{n!}{(k!)^i (n - ki)!} \sum_{j=k+1}^{n-ki} s_{j,n-ki} + [n/k] \frac{c_k^{n/k}}{(n/k)!} \frac{n!}{(k!)^{n/k}} \tag{5}$$

$$= \sum_{i=1}^{\lfloor n/k \rfloor} \frac{n!}{k^i i! (n - ki)!} \sum_{j=k+1}^{n-ki} s_{j,n-ki} + [n/k] \frac{n!}{(n/k)! k^{n/k}}. \tag{6}$$

In order to simplify the notation from [12] to fit our purpose here, we change slightly the notation from $L_{k,n}^s$ to $s_{k,n}$.

For a fixed n , we have the following properties:

$$s_{n,n} = (n - 1)!, \quad s_{k,n} = 0 \text{ for } \lfloor n/2 \rfloor + 1 \leq k \leq n - 1, \quad \text{and} \quad \sum_{k=1}^n s_{k,n} = n!.$$

We have for a fixed $n \geq 1$ that

$$\mathbf{P}\{X_n = k\} = \frac{s_{k,n}}{n!} \quad \text{for } 1 \leq k \leq n.$$

In Section 2, we evaluate C from (3) using different approaches. Another of our goals in Section 3, is to evaluate the generalized Buchstab¹ function with parameter $K > 0$ defined by

$$\Omega_K(x) = \begin{cases} 1 & \text{for } 1 \leq x < 2, \\ 1 + K \int_2^x \frac{\Omega_K(u-1)}{u-1} du & \text{for } x \geq 2. \end{cases} \tag{7}$$

The quantity $1/\Omega_K(x)$ gives the fraction of n -objects with large smallest components; more precisely, Theorem 1.1 from [1] stipulates that

$$\lim_{n \rightarrow \infty} \frac{s_{\lfloor xn \rfloor, \lfloor xn \rfloor}}{\sum_{i=n}^{\lfloor xn \rfloor} s_{\lfloor xn \rfloor, i}} = \frac{1}{\Omega_K(x)} \quad \text{for } x > 1.$$

For the sake of completeness and to gain insight into how the Buchstab function connects to combinatorial analysis, we end this introduction by briefly recalling how Buchstab introduced his function ω when studying the factorization of natural numbers into primes. The primes are like the irreducible factors of a polynomial, or the cycles of a permutation, etc. Let $\xi \in \{1, \dots, n\}$ with its decomposition into

¹We thank an anonymous referee for bringing to our attention that the function considered here is not a possible generalization of the original Buchstab because there is no K such that Ω_K coincides with ω on the interval $[1, 2)$.

primes given as $p_1(\xi) \cdots p_k(\xi) = \xi$ such that $p_1(\xi) \leq p_2(\xi) \leq \dots \leq p_k(\xi)$. We count the number of ξ 's with their smallest prime factor less than m ; in other words, set

$$\Psi(n, m) = \text{card}\{\xi \in \{1, \dots, n\} : p_1(\xi) \leq m\}.$$

Then in [3] it is shown that

$$\Psi(n, m) = 1 + \sum_{p \leq m} \Psi\left(\frac{n}{p}, p\right) \quad \text{for all } 1 < m \leq n.$$

The previous summation is over all the primes p less than or equal to m . The functional equation given by Ψ is related to the Dickman function, which we do not discuss here; see [15] for a detailed analysis of the Dickman function and the Buchstab function.

2. Approaches

2.1. Analytic Estimation

This section mostly recalls results from [11] and [13]. The approach from [13] to obtain the limiting quantities for $\mathbf{P}\{X_n \geq k\}$ and $\mathbf{E}(X_n^\ell)$ as $k, n \rightarrow \infty$ and $\ell \geq 1$ uses singularity analysis of exponential generating functions for combinatorial objects. For an in-depth coverage of singularity analysis applied to combinatorics, see [6].

Permutations form a typical class of combinatorial objects that we choose here for our discussion, but the results are not limited only to permutations. The cycles are the irreducible components of a permutation. Let $C(z) = \sum_{i=0}^\infty C_i z^i / i!$ be the exponential generating function for counting cycles of given lengths. Then, the exponential generating function for counting permutations of given sizes is

$$L(z) = \exp(C(z)) = \sum_{i=0}^\infty L_i \frac{z^i}{i!}.$$

For a fixed $n > 0$, we are interested in counting permutations with the smallest cycles of length at least k for $1 \leq k \leq n$. Let $S(z)$ be the generating function for counting permutations with the smallest cycles of length at least k for $1 \leq k \leq n$. Then we have

$$S(z) = \exp\left(\sum_{i=k}^\infty C_i \frac{z^i}{i!}\right) - 1 = \sum_{i=0}^\infty S_i \frac{z^i}{i!}.$$

Therefore, the tail of the probability distribution of X_n is given by

$$\mathbf{P}\{X_n \geq k\} = \frac{S_n}{L_n}.$$

Using singularity analysis, in [13] it is shown that if $k, n \rightarrow \infty$, then

$$\mathbf{P}\{X_n \geq k\} = \frac{1}{k} \omega\left(\frac{n}{k}\right) + O\left(\frac{1}{k^{1+\epsilon}}\right) \quad \text{for some } \epsilon > 0. \tag{8}$$

Theorem 1 states the asymptotic behavior of the moments.

Theorem 1. *For some function $h(n)$ that tends to infinity slower than $\log(n)$ and for some $\epsilon > 0$ independent of n , we have that*

$$\begin{aligned} \mathbf{E}(X_n) &= e^{-\gamma} \log(n) \left(1 + O\left(\frac{h(n)}{\log(n)}\right)\right), \\ \mathbf{E}(X_n^\ell) &= \ell n^{\ell-1} \left(\int_1^\infty \frac{\omega(x)}{x^\ell} dx\right) \left(1 + O\left(\frac{1}{n^\epsilon}\right)\right) \quad \text{for integers } \ell \geq 2. \end{aligned}$$

Proof. We consider the case when $\ell \geq 2$. We give the main steps for proving Theorem 1. By definition, we have

$$\mathbf{E}(X_n^\ell) = \sum_{k=1}^\infty (k^\ell - (k-1)^\ell) \mathbf{P}\{X_n \geq k\}. \tag{9}$$

Let $\nu(n) = \lfloor n^{\epsilon'} \rfloor$ such that $0 < \epsilon' < \epsilon$ where ϵ is given from (8). Then $\nu(n) \rightarrow \infty$ as $n \rightarrow \infty$, so we split the sum from (9) using ν , and we obtain

$$\begin{aligned} \mathbf{E}(X_n^\ell) &= \sum_{k=1}^{\nu(n)-1} (k^\ell - (k-1)^\ell) \mathbf{P}\{X_n \geq k\} + \sum_{k=\nu(n)}^\infty (k^\ell - (k-1)^\ell) \mathbf{P}\{X_n \geq k\} \\ &\stackrel{\text{def}}{=} S_1 + S_2. \end{aligned}$$

Using (8), and the fact that $\mathbf{P}\{X_n \geq n+1\} = 0$, we have $S_1 = O((\nu(n))^{\ell-1})$ because $(k^\ell - (k-1)^\ell) \in O(k^{\ell-1})$ and $\mathbf{P}\{X_n \geq k\} \in O(1/k^{1+\epsilon})$ in the range $1 \leq k < \nu(n)$. In the range $\nu(n) \leq k \leq n$, we have $(k^\ell - (k-1)^\ell) \in O(\ell k^{\ell-1})$, and therefore

$$\begin{aligned} S_2 &= \sum_{k=\nu(n)}^\infty (k^\ell - (k-1)^\ell) \mathbf{P}\{X_n \geq k\} \\ &= \ell \left(\sum_{k=\nu(n)}^n k^{\ell-2} \omega\left(\frac{n}{k}\right) \right) (1 + O(\nu(n)^{-\epsilon})). \end{aligned} \tag{10}$$

The sum within (10) is a Riemann sum estimated by its corresponding integral

$$\begin{aligned} \sum_{k=\nu(n)}^n k^{\ell-2} \omega\left(\frac{n}{k}\right) &= \int_0^n t^{\ell-2} \omega\left(\frac{n}{t}\right) dt + O\left(\frac{1}{n}\right) \\ &= n^{\ell-1} \int_1^\infty \frac{\omega(x)}{x^\ell} dx + O\left(\frac{1}{n}\right) \quad \text{with } \frac{n}{t} = x. \end{aligned}$$

The proof for the case $\ell = 1$ is quite similar, and the range $\nu(n) \leq k \leq n$ is divided further into two ranges $\nu(n) \leq k < n\mu(u)$ and $n\mu(n) \leq k \leq n$ where $\mu(n)$, for some well-chosen function μ , is defined as in [11]. \square

Remark 2. The sum in (10) goes up to n inclusively and not $n/2$; thus, the range of integration starts at 1 and not 2. Because $\mathbf{P}\{X_n = k\} = 0$ for $\lfloor n/2 \rfloor + 1 \leq k \leq n-1$, we also point out that

$$\mathbf{P}\{X_n \geq k\} = \sum_{i=k}^n \mathbf{P}\{X_n = i\} = \mathbf{P}\{X_n = n\} \quad \text{for } \lfloor n/2 \rfloor + 1 \leq k \leq n.$$

Going back to the variance of X_n , we have the following theorem that ends our section on the analytical estimation for $\text{Var}(X_n)/n$ as $n \rightarrow \infty$.

Theorem 2. For some $\epsilon > 0$ independent of n , we have that

$$\text{Var}(X_n) = nC \left(1 + O\left(\frac{1}{n^\epsilon}\right) \right) \quad \text{with } C = 2 \int_1^\infty \frac{\omega(x)}{x^2} dx.$$

Proof. We have by definition that $\text{Var}(X_n) = \mathbf{E}(X_n^2) - (\mathbf{E}(X_n))^2$. We use (8) and consider the second moment. Hence we have

$$\begin{aligned} \mathbf{E}(X_n^2) &= \sum_{k=1}^\infty (k^2 - (k-1)^2) \mathbf{P}\{X_n \geq k\} = \sum_{k=1}^\infty (2k-1) \mathbf{P}\{X_n \geq k\} \\ &= \sum_{k=1}^n (2k-1) \mathbf{P}\{X_n \geq k\} \\ &= \sum_{k=1}^n (2k-1) \left(\frac{1}{k} \omega\left(\frac{n}{k}\right) + O\left(\frac{1}{k^{1+\epsilon}}\right) \right) \quad \text{for some } \epsilon > 0 \\ &\sim 2 \sum_{k=1}^n \omega\left(\frac{n}{k}\right). \end{aligned} \tag{11}$$

The expression (11) is a Riemann sum estimated in a way similar to that of Theorem 1. The quantity $(\mathbf{E}(X_n))^2$ is negligible compared to $\mathbf{E}(X_n^2)$ as $n \rightarrow \infty$. Hence, we have that

$$\text{Var}(X_n) \sim 2n \int_1^\infty \frac{\omega(x)}{x^2} dx \quad \text{as } n \rightarrow \infty.$$

Reference [14] proves that $\omega(x) \rightarrow e^{-\gamma}$ where γ is the Euler-Mascheroni constant. More specifically, it proves that $|\omega(x) - e^{-\gamma}| < 10^{-4}$ for $x > 4$. Therefore, we have that

$$C = 2 \int_1^\infty \frac{\omega(x)}{x^2} dx = 2 \int_1^4 \frac{\omega(x)}{x^2} dx + 2 \int_4^\infty \frac{e^{-\gamma}}{x^2} dx + 2 \int_4^\infty \frac{\omega(x) - e^{-\gamma}}{x^2} dx.$$

Using the quantities from [11] for

$$2 \int_2^\infty \frac{\omega(x)}{x^2} dx = 0.5586\dots,$$

and, this time, taking into account the evaluation of the integral over $[1, 2]$ that yields exactly $3/4$, we obtain up to four significant figures that $C = 1.3068\dots$, and thus

$$\frac{\text{Var}(X_n)}{n} \rightarrow 1.3068\dots \quad \text{as } n \rightarrow \infty.$$

The proof is now complete. □

2.2. Numerical Integration

We use an idea from [8] in Theorem 3 to evaluate $\omega(x)$, for any $x \geq 1$, with arbitrary finite precision. We use Theorem 3 to evaluate C . The quantity n in this section is not the same as previously stated, which stood for the number of elements considered in our combinatorial object, while n here stands for the integral part of a real number, as is standard in numerical approximations.

We recall that we need to evaluate

$$C = 2 \int_1^\infty \frac{\omega(t)}{t^2} dt = \lim_{n \rightarrow \infty} \frac{\text{Var}(X_n)}{n}. \tag{12}$$

For notational simplicity, we use $f : [1, \infty) \rightarrow [0, 1]$ to denote the function $x \mapsto \omega(x)/x^2$. As mentioned previously, $|\omega(x) - e^{-\gamma}| < 10^{-4}$ for $x > 4$, and f is therefore bounded. The function f is also continuous because it is the composition of two continuous functions on $[1, \infty)$. We have that $f(x) \rightarrow 0$ as $x \rightarrow \infty$. Hence, the Riemann sum of f is convergent. We can approximate numerically its Riemann sum, that is, $\int_1^\infty f(t)dt$, up to a desired accuracy by truncating the integral; this is because $f(x) \rightarrow 0$.

A popular method to approximate an integral is the trapezoidal method with a regular grid of points. Consider the interval $[1, n^*]$ where $n^* \in \mathbb{N}$ shall be determined later. Given the nature of ω (and so f), we consider for now an interval of the form $[n, n + 1]$ where $n \in \mathbb{N}$. A point from a regular grid on $[n, n + 1]$ can be put conveniently into the form $x_i = n + i\delta$ for $0 \leq i \leq \ell$ where $\delta = 2^{-\ell}$. We therefore have that

$$\sum_{i=0}^{2^\ell-1} \delta \frac{(f(n + i\delta) + f(n + (i + 1)\delta))}{2} \rightarrow \int_n^{n+1} f(t)dt \quad \text{as } \ell \rightarrow \infty. \tag{13}$$

To evaluate C with four significant digits, we can select $n^* = 10000$ and $\ell = 14$ so that $\delta < 10^{-4}$ using, for instance, the sharp bounds on numerical integration from [4]. Now, it remains to know how to compute numerically $\omega(x)$ for $x \geq 1$, which we do using the Taylor series as given by Theorem 3.

Theorem 3. Consider the Taylor expansions of ω with respect to the z variable for each unit length interval of the form $[n, n + 1)$. More precisely let

$$\omega\left(n + \frac{1+z}{2}\right) = \sum_{i=0}^{\infty} c_{n,i} z^i \quad \text{for } n \geq 1 \text{ and for } -1 \leq z < 1.$$

Let $c_{n,i}$ the i -th term for the n -th sequence \mathbf{c}_n for $n \geq 1$ and $i \geq 0$. Then we have

$$\begin{aligned} c_{1,i} &= \frac{2}{3} \left(\frac{-1}{3}\right)^i, \\ c_{n+1,0} &= \frac{1}{2n+3} \sum_{i=0}^{\infty} c_{n,i} \left(2(n+1) + \frac{(-1)^i}{i+1}\right) \quad \text{for } n > 1, \\ c_{n+1,i} &= \frac{1}{2n+3} \left(\frac{c_{n,i}}{n} - c_{n+1,i-1}\right) \quad \text{for } n > 1 \text{ and } i \geq 1. \end{aligned}$$

Proof. Let $n \geq 1$ and let $x = n + t \geq 1$ with $n = \lfloor x \rfloor$ and $0 \leq t < 1$. If ω has a Taylor expansion in $[n, n + 1)$, that is, it has the coefficients $c_{n,i}$, then we obtain the coefficients $c_{n+1,i}$ of the Taylor expansion in $[n + 1, n + 2)$ as follows. We integrate the difference-differential equation (2) and have that

$$\begin{aligned} \int_{u=n+1}^{u=n+1+t} d(u\omega(u)) &= (n+1+t)\omega(n+1+t) - (n+1)\omega(n+1) \\ &= \int_{u=n+1}^{u=n+1+t} \omega(u-1) du \\ &= \int_{x=0}^{x=t} \omega(n+x) dx, \quad \text{with } u = n+1+x. \end{aligned}$$

The affine transformation $t = z = 2t + 1$ transforms the fractional part $t \in [0, 1)$ into a centered-around-0 value $z \in [-1, 1)$. Equivalently, $t = (z + 1)/2$, and we have that

$$\begin{aligned} \left(n+1 + \frac{z+1}{2}\right)\omega\left(n+1 + \frac{z+1}{2}\right) - (n+1)\omega(n+1) & \tag{14} \\ &= \int_{v=0}^{v=(z+1)/2} \omega(n+1+v) dv \\ &= \frac{1}{2} \int_{u=-1}^{u=z} \omega\left(n + \frac{u+1}{2}\right) du \quad \text{with } v = \frac{u+1}{2}. \tag{15} \end{aligned}$$

Using the Taylor expansion around $u = 0$ of ω in the interval $[n, n + 1)$ in terms of the dummy variable of integration, we have

$$\omega\left(n + \frac{u+1}{2}\right) = \sum_{i=0}^{\infty} c_{n,i} u^i \quad \text{for } -1 \leq u \leq z < 1. \tag{16}$$

Hence by substituting (16) into (15):

$$\int_{u=-1}^{u=z} \omega\left(n + \frac{u+1}{2}\right) du = \int_{-1}^z \sum_{i=0}^{\infty} c_{n,i} u^i du = \sum_{i=0}^{\infty} c_{n,i} \frac{(z^{i+1} - (-1)^{i+1})}{i+1}. \tag{17}$$

By continuity of ω , we have also that

$$\lim_{z \rightarrow 1} \omega\left(n + \frac{z+1}{2}\right) = \omega(n+1) = \lim_{z \rightarrow 1} \sum_{i=0}^{\infty} c_{n,i} z^i = \sum_{i=0}^{\infty} c_{n,i}. \tag{18}$$

Using the Taylor expansion around $z = 0$ of ω in the interval $[n+1, n+2)$, we obtain

$$\omega\left(n+1 + \frac{z+1}{2}\right) = \sum_{i=0}^{\infty} c_{n+1,i} z^i \quad \text{for } -1 \leq z < 1.$$

Then substituting (18) into (14) and (17) into (15) yields:

$$(2n+3+z) \sum_{i=0}^{\infty} c_{n+1,i} z^i = 2(n+1) \sum_{i=0}^{\infty} c_{n,i} + \sum_{i=0}^{\infty} c_{n,i} \frac{(z^{i+1} - (-1)^{i+1})}{i+1}. \tag{19}$$

Substituting $z = 0$ in (19), we get

$$c_{n+1,0} = \frac{1}{2n+3} \sum_{i=0}^{\infty} c_{n,i} \left(2(n+1) + \frac{(-1)^i}{i+1}\right). \tag{20}$$

By using (20) and equating coefficients with the same power of z , we find $c_{n+1,i}$ for $i \geq 1$:

$$\begin{aligned} (2n+3+z)c_{n+1,0} + (2n+3+z) \sum_{i=1}^{\infty} c_{n+1,i} z^i \\ = 2(n+1) \sum_{i=0}^{\infty} c_{n,i} + \sum_{i=0}^{\infty} c_{n,i} \frac{(z^{i+1} + (-1)^i)}{i+1}, \end{aligned}$$

$$c_{n+1,0}z + (2n+3+z) \sum_{i=1}^{\infty} c_{n+1,i} z^i = c_{n,0}z + \sum_{i=1}^{\infty} c_{n,i} \frac{z^{i+1}}{i+1},$$

and

$$\begin{aligned} (2n+3+z) \sum_{i=1}^{\infty} c_{n+1,i} z^i \\ = (2n+3)c_{n+1,1}z + (2n+3) \sum_{i=2}^{\infty} c_{n+1,i} z^i + \sum_{i=1}^{\infty} c_{n+1,i} z^{i+1}. \end{aligned}$$

The previous equation holds if and only if

$$((2n + 3)c_{n+1,i} + c_{n+1,i-1})z^i = \frac{c_{n,i-1}z^i}{i} \quad \text{for all } i \geq 1.$$

We finally find the Taylor expansion of $1/x$ around $x = 1$ with $1 \leq x = 1 + t \leq 2$ and $t = (1 + z)/2$ for $-1 \leq z < 1$, and have

$$\omega\left(1 + \frac{1+z}{2}\right) = \frac{2}{3} \frac{1}{(1+(z/3))} = \frac{2}{3} \sum_{i=0}^{\infty} \left(\frac{-1}{3}\right)^i z^i = \sum_{i=0}^{\infty} c_{1,i} z^i.$$

The proof is now complete. □

We point out that the centered-around-0 flavor of the Taylor expansion with coefficients \mathbf{c}_n allows faster convergence around the endpoints n and $n + 1$; see [8]. We compute the first n^* sequences with their first J terms, provided we have a library that does real arithmetic with finite and arbitrary precision.

Algorithm 1 Trapezoidal rule by using Taylor coefficient of the Buchstab function on the interval $[n, n + 1)$ for $n \in \mathbb{N}$

Input: $\ell, n, \{c_{n,j}\}_{j=0}^J$
Output: s , the sum from (13).
 1: $\delta \leftarrow 2^{-\ell}$
 2: $s \leftarrow 0$
 3: **for** $i = 0$ **to** $2^\ell - 1$ **do**
 4: $y_0 \leftarrow 0$
 5: $y_1 \leftarrow 1$
 6: $t_0 \leftarrow i\delta$
 7: $t_1 \leftarrow (i + 1)\delta$
 8: $z_0 \leftarrow 1$
 9: $z_1 \leftarrow 1$
 10: **for** $j = 0$ **to** J **do**
 11: $y_0 \leftarrow y_0 + c_{n,j}z_0$
 12: $y_1 \leftarrow y_1 + c_{n,j}z_1$
 13: $z_0 \leftarrow z_0(2t_0 - 1)$
 14: $z_1 \leftarrow z_1(2t_1 - 1)$
 15: **end for**
 16: $s \leftarrow s + \frac{y_0}{(n+t_0)^2} + \frac{y_1}{(n+t_1)^2}$
 17: **end for**
 18: $s \leftarrow \frac{s\delta}{2}$

To obtain C , we iteratively call Algorithm 1 for values of $n = 1, 2, \dots, n^*$ with the coefficients for the Taylor expansion of ω on the interval $[n, n + 1)$. We add the result

of all iterations together and obtain $C = 1.3070\dots$, which confirms comfortably the estimation from Section 2.1.

We end this section with comments about Algorithm 1. We have from line (7) that $t_1 = t_0 + \delta$. The loop at line (10) computes the Taylor polynomial of degree J of the Buchstab function $\omega(n + (1 + z)/2)$ for the specific values of $z = z_0$ and $z = z_1$. During the j -th iteration at lines (11) and (12), we have that $y_b = \sum_{k=0}^j c_{n,k} z_b^k$ for $b = 0$ and $b = 1$, respectively. Lines (13) and (14) are for updating z_0 and z_1 , respectively, for the next iteration, that is, the $(j + 1)$ -th iteration. We recall the meaning of the left side of the limiting expression in Equation (13) is that the height of a rectangle is $(f(n + i\delta) + f(n + (i + 1)\delta))/2$ with $f(x) = \omega(x)/x^2$ in our case, and its length is δ ; therefore, line (16) sums over the heights of all the rectangles. Averaging two consecutive heights by 2 is carried out only once at line (18) to save a few operations. Similarly, we take into account the length δ , which is identical for each rectangle, only once at line (18).

2.3. Recurrence Relation

We compute the probability distribution of X_n and then compute $\text{Var}(X_n)$ for values of $n = 1, 2, \dots, 4000$. Recalling (1), we have that

$$\text{Var}(X_n) = \sum_{k=1}^n (k - \mathbf{E}(X_n))^2 \mathbf{P}\{X_n = k\}.$$

Because

$$\mathbf{E}(X_n) = \sum_{k=1}^n k \mathbf{P}\{X_n = k\} \quad \text{and} \quad \mathbf{P}\{X_n = k\} = \frac{s_{k,n}}{n!},$$

the variance is a rational number, which is suitable to control the accuracy, as follows:

$$\frac{n! \sum_{k=1}^n k^2 s_{n,k} - \left(\sum_{k=1}^n k s_{n,k}\right)^2}{(n!)^2}.$$

We divide the quantity $\text{Var}(X_n)$ by n in order to normalize it. We recall that $\text{Var}(X_n) = C(n + O(n^{-\epsilon}))$ for some $\epsilon > 0$. When computing exactly $\text{Var}(X_n)$ for a fixed n and comparing with the asymptotic formula, one would need the hidden factor of $n^{-\epsilon}$ and the value ϵ itself in order make a fair comparison; we nevertheless obtain numbers that are very close to the numbers from Sections 2.1 and 2.2:

$$\begin{aligned} \frac{\text{Var}(X_{1000})}{1000} &= 1.3004\dots, & \frac{\text{Var}(X_{2000})}{2000} &= 1.3036\dots, \\ \frac{\text{Var}(X_{3000})}{3000} &= 1.3047\dots, & \frac{\text{Var}(X_{4000})}{4000} &= 1.3053\dots \end{aligned}$$

The memory size on the machines available to us is the main limitation; however, it is enough to assert C up to two significant digits. A space of $12.7GB$ is needed to compute the triangular table for $n = 4000$. Storing the values in a triangular array allows us to compute the recurrence relation easily. Trimming the array of potentially unused cells is tough as n grows. Each array cell holds $s_{n,k}$ for a pair (n, k) . The values $s_{n,k}$ are given by (6). We could compress the array slightly for $s_{n,k}$ when $\lfloor n/2 \rfloor + 1 \leq k \leq n - 1$ using methods described in [10] for instance, but we would not gain much for large values of n (like $n > 1000$) in terms of space and would yield a more complicated code. A possible algorithm for counting the $s_{n,k}$ is as in Algorithm 2.

Algorithm 2 Computing $s_{n,k}$

Input: N

Output: $s_{n,k}$ for $1 \leq n \leq N$ and $1 \leq k \leq n$

```

1:  $s_{0,0} \leftarrow 1$ 
2: for  $n = 1$  to  $N$  do
3:    $s_{n,0} \leftarrow 0$ 
4:    $s_{n,n} \leftarrow (n - 1)!$ 
5: end for
6: for  $n = 2$  to  $N$  do
7:   for  $k = 1$  to  $\lfloor n/2 \rfloor$  do
8:      $t_1 \leftarrow 0$ 
9:     for  $i = 1$  to  $\lfloor n/k \rfloor$  do
10:       $u_1 \leftarrow 0$ 
11:      for  $j = k + 1$  to  $n - ki$  do
12:         $u_1 \leftarrow u_1 + s_{n-ki,j}$ 
13:      end for
14:      if  $k + 1 \leq n - ki$  then
15:         $u_1 \leftarrow u_1 \frac{n!}{i!k^i(n-ki)!}$ 
16:      end if
17:       $t_1 \leftarrow t_1 + u_1$ 
18:    end for
19:     $t_2 \leftarrow 0$ 
20:    if  $k$  divides  $n$  then
21:       $t_2 \leftarrow \frac{n!}{(n/k)!k^{n/k}}$ 
22:    end if
23:     $s_{n,k} \leftarrow t_1 + t_2$ 
24:  end for
25: end for

```

We make just a few comments about Algorithm 2. From a data structure point of view, $n = 0$ and $k = 0$ are boundaries for the array, and lines (1) and (3) define

the programming boundaries but are not part of the combinatorial objects and their related probability distributions a fortiori. The loop at line (7) runs up to $\lfloor n/2 \rfloor$ because we assume that $s_{n,k}$ are initialized to 0 by default for all valid n and k ; this is usually the case in most advanced programming languages when declaring data structures.

We end this section with a small example. Table 1 shows $s_{n,k}$ for $1 \leq n \leq 10$.

n	k									
	1	2	3	4	5	6	7	8	9	10
10	2293839	525105	223200	151200	72576	0	0	0	0	362880
9	229384	52632	22400	18144	0	0	0	0	40320	
8	25487	5845	2688	1260	0	0	0	5040		
7	3186	714	420	0	0	0	720			
6	455	105	40	0	0	120				
5	76	20	0	0	24					
4	15	3	0	6						
3	4	0	2							
2	1	1								
1	1									

Table 1: Values of $s_{n,k}$ for $1 \leq n \leq 10$.

3. Generalized Buchstab Function

We recall the definition of the generalized Buchstab function with parameter $K > 0$, which is

$$\Omega_K(x) = \begin{cases} 1 & \text{for } 1 \leq x < 2, \\ 1 + K \int_2^x \frac{\Omega_K(u-1)}{u-1} du & \text{for } x \geq 2. \end{cases} \tag{21}$$

The values of $1/\Omega_K(x)$ are asymptotic proportions of the large smallest components, as proved in [1]. More precisely, we recall that $s_{n,k}$, given as in (5) of Section 1, is the number of combinatorial n -objects with their smallest components having length k . For instance, classes of objects with parameter $K = 1/2$ include 2-regular graphs, surjective maps, etc. Classes of objects with parameter $K = 1$ include derangements, permutations, monic polynomials over a finite field, etc. The quantity $\sum_{i=k}^n s_{n,i}$ is the number of n -objects for which the smallest component has a size of at least k for $1 \leq k \leq n$. Let $x > 1$ and consider the ratio

$$\frac{s_{\lfloor xn \rfloor, \lfloor xn \rfloor}}{\sum_{i=n}^{\lfloor xn \rfloor} s_{\lfloor xn \rfloor, i}}. \tag{22}$$

Then it is shown in [1] that, for $x > 1$,

$$\lim_{n \rightarrow \infty} \frac{s_{\lfloor xn \rfloor, \lfloor xn \rfloor}}{\sum_{i=n}^{\lfloor xn \rfloor} s_{\lfloor xn \rfloor, i}} = \frac{1}{\Omega_K(x)}. \tag{23}$$

The limiting quantity (23) justifies our interest in evaluating the generalized Buchstab function.

We remark that from now on and up to Table 2 inclusively, the symbol n no longer refers to the size of a combinatorial object.

Following the ideas exposed in Section 2.2, let $n \geq 1$ be a natural number, and let $c_{n,i}$ be i -th coefficient of the Taylor expansion for $\Omega_K(z)$ in the interval $[n, n + 1)$ with $-1 \leq z < 1$. More precisely, let

$$\Omega_K\left(n + \frac{1+z}{2}\right) = \sum_{i=0}^{\infty} c_{n,i} z^i \quad \text{for } -1 \leq z < 1. \tag{24}$$

As expected, the sequence $(c_{n,i})_{i \geq 0}$ depends on the previous sequence $(c_{n-1,i})_{i \geq 0}$ for $n > 2$. Our library can compute over \mathbb{R} with arbitrary finite precision. The variable z in (24) is the fractional part of $x \in [n, n + 1)$ centered around 0.

Theorem 4. *For $K > 0$, consider the Taylor expansions of Ω_K with respect to the z variable for each unit length interval of the form $[n, n + 1)$. More precisely, let*

$$\Omega_K\left(n + \frac{1+z}{2}\right) = \sum_{i=0}^{\infty} c_{n,i} z^i \quad \text{for } n \geq 1 \text{ and for } -1 \leq z < 1.$$

For $n \geq 1$ and $i \geq 0$, and let α_i be defined by

$$\alpha_i = \sum_{j=0}^i \frac{(-1)^{i-j}}{(2n-1)^{i-j}} c_{n-1,j}.$$

Then we have

$$\begin{aligned} c_{1,0} &= 1, \\ c_{1,i} &= 0 \quad \text{for } i \geq 1, \\ c_{2,0} = c_{2,0} &= 1 + K \sum_{i=1}^{\infty} \frac{1}{i2^i}, \\ c_{2,i} &= K \sum_{j=i}^{\infty} \frac{(-1)^{j-1}}{j2^j} \binom{j}{i} \quad \text{for } i \geq 1, \\ c_{n,0} &= \sum_{i=0}^{\infty} c_{n-1,i} - \frac{K}{2n-1} \sum_{i=0}^{\infty} \frac{(-1)^{i+1} \alpha_i}{i+1} \quad \text{for } n \geq 3, \\ c_{n,i} &= \frac{K \alpha_{i-1}}{(2n-1)i} \quad \text{for } n \geq 3 \text{ and } i \geq 1. \end{aligned}$$

Proof. For $x \in [1, 2)$, the function Ω_K is constant and so $c_{1,0} = 1$ and $c_{1,i} = 0$ for $i \geq 1$.

For $2 \leq x = 2 + ((1 + z)/2) < 3$, the coefficients of the Taylor expansion are $1 + K \log(2 + (1 + z)/2)$; hence the coefficients are given by

$$c_{2,0} = 1 + K \sum_{j=1}^{\infty} \frac{1}{j2^j} \quad \text{and} \quad c_{2,i} = K \sum_{j=i}^{\infty} \frac{(-1)^{j-1}}{j2^j} \binom{j}{i} \quad \text{for } i \geq 1. \quad (25)$$

Given $x \geq 3$ such that $x = n + ((z + 1)/2)$ and $n \geq 3$, we assume that the sequence $(c_{n-1,i})_{i \geq 0}$ is known. We have

$$\begin{aligned} \Omega_K\left(n + \left(\frac{1+z}{2}\right)\right) &= \sum_{i=0}^{\infty} c_{n,i} z^i \\ &= 1 + K \int_2^{n+(1+z)/2} \frac{\Omega_K(u-1)}{u-1} du \\ &= 1 + K \int_2^n \frac{\Omega_K(u-1)}{u-1} du + K \int_n^{n+(1+z)/2} \frac{\Omega_K(u-1)}{u-1} du \\ &= \Omega_K(n) + K \int_{u=n}^{u=n+(1+z)/2} \frac{\Omega_K(u-1)}{u-1} du \\ &= \Omega_K(n) + K \int_{v=-1}^{v=z} \frac{\Omega_K(n-1+(v+1)/2)}{2n-1+v} dv \quad \text{with } v = 2u - 2n - 1 \\ &= \Omega_K(n) + \frac{K}{2n-1} \int_{u=-1}^{u=z} \left(\sum_{i=0}^{\infty} c_{n-1,i} u^i\right) \left(\sum_{i=0}^{\infty} \frac{(-1)^i u^i}{(2n-1)^i}\right) du \\ &= \Omega_K(n) + \frac{K}{2n-1} \int_{u=-1}^{u=z} \sum_{i=0}^{\infty} \left(\sum_{j=0}^i \frac{(-1)^{i-j}}{(2n-1)^{i-j}} c_{n-1,j}\right) u^i du \\ &= \Omega_K(n) + \frac{K}{2n-1} \int_{u=-1}^{u=z} \sum_{i=0}^{\infty} \alpha_i u^i du \\ &= \Omega_K(n) - \frac{K}{2n-1} \sum_{i=0}^{\infty} \frac{(-1)^{i+1} \alpha_i}{i+1} + \frac{K}{2n-1} \sum_{i=0}^{\infty} \frac{\alpha_i z^{i+1}}{i+1}. \end{aligned} \quad (26)$$

The continuity of Ω_K implies that

$$\Omega_K(n) = \lim_{z \rightarrow 1} \Omega_K\left(n - 1 + \frac{1+z}{2}\right) = \lim_{z \rightarrow 1} \sum_{i=0}^{\infty} c_{n-1,i} z^i = \sum_{i=0}^{\infty} c_{n-1,i}.$$

Hence (26) is rewritten as

$$\Omega_K\left(n + \frac{1+z}{2}\right) = \sum_{i=0}^{\infty} c_{n-1,i} - \frac{K}{2n-1} \sum_{i=0}^{\infty} \frac{\alpha_i (-1)^{i+1}}{i+1} + \frac{K}{2n-1} \sum_{i=0}^{\infty} \frac{\alpha_i z^{i+1}}{i+1}$$

$$= c_{n,0} + \sum_{i=1}^{\infty} \frac{K\alpha_{i-1}}{(2n-1)^i} z^i = c_{n,0} + \sum_{i=1}^{\infty} c_{n,i} z^i.$$

The proof is complete. □

For instance, by reading $\Omega_1(2^{13})$ from the left half of Table 2 and recalling (22), the proportion of random permutations on at least 2^{14} elements, and with a cycle of smallest length at least 2^{13} , is close to $1/\Omega_1(2^{13}) \approx 0.000218$. We note that if the number of permuted elements is precisely 2^{14} , then there will be no smallest component of size at least 2^{13} ; one can observe this from the recurrence relation in Section 2.3 as well.

Similarly, by reading $\Omega_{1/2}(2^{13})$ from the right half of Table 2 and recalling (22), the proportion of random 2-regular graphs with at least 2^{14} vertices, and with a large smallest component of at least 2^{13} , is close to $1/\Omega_{1/2}(2^{13}) \approx 0.0131$. We note that if the number of vertices is exactly 2^{14} , then there will be no smallest component of size at least 2^{13} .

$K = 1$				$K = 1/2$			
x	$\Omega_K(x)$	x	$\Omega_K(x)$	x	$\Omega_K(x)$	x	$\Omega_K(x)$
1	1	16	8.9874	1	1	16	3.3302
2	1	32	17.9749	2	1	32	4.7470
3	1.6941	64	35.9498	3	1.3470	64	6.7397
4	2.2468	128	71.8997	4	1.5866	128	9.5501
5	2.8085	256	143.7995	5	1.7971	256	13.5191
6	3.3703	512	287.5991	6	1.9856	512	19.1282
7	3.9320	1024	575.1983	7	2.1579	1024	27.0580
8	4.4937	2048	1150.3966	8	2.3175	2048	38.2705
9	5.0554	4096	2300.7932	9	2.4669	4096	54.1260
10	5.6171	8192	4567.8834	10	2.6077	8192	76.5480

Table 2: A few values of $\Omega_K(x)$ for $K = 1$ and $K = 1/2$.

We conclude this section by mentioning that [5] gives values for $1/\Omega_K(x)$ with $x = 2, 3, 4, 5$ and that, if we invert values from Table 2 for $x = 2, 3, 4, 5$, they agree with those from [5].

4. Conclusion

In this paper, we computed the normalization constant of the variance of the distribution of the smallest component of random combinatorial objects. We used different approaches: an analytic method based on the singularity analysis for gen-

erating functions, a numerical integration method using Taylor expansions for the Buchstab function, and the recurrence relation for counting the number of smallest components. All the methods yield $1.3070\dots$. We also showed how to compute the value of the generalized Buchstab function by recursively building sequences of Taylor expansions for each unit interval of the form $[n, n + 1)$ where $n \in \mathbb{N} \setminus \{0\}$. We can compute the asymptotic proportion of large smallest components for various random combinatorial objects by obtaining very accurate values of the generalized Buchstab function.

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